

Supplemental Material S1. Analysis code.

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First let's load in the data.

```
library(tidyverse)
#> Warning: package 'tibble' was built under R version 4.0.3
model_data <- drake::readadd("data_paper_model_counts") %>%
  filter(aligner != "Manual")

# See a few values from each column
glimpse(model_data)
#> Rows: 840
#> Columns: 10
#> $ class      <chr> "fricatives", "fricatives", "fricatives", "fricatives"...
#> $ aligner    <fct> Kaldi, Kaldi, Kaldi, Kaldi, Kaldi, Kaldi, Kaldi, Kaldi...
#> $ child      <chr> "c01", "c02", "c03", "c04", "c05", "c06", "c07", "c08"...
#> $ age        <dbl> 50, 39, 41, 76, 43, 75, 50, 60, 55, 48, 57, 77, 78, 70...
#> $ age_scale  <dbl> -0.83333333, -1.75000000, -1.58333333, 1.33333333, -1....
#> $ n_files    <int> 67, 37, 38, 68, 28, 67, 47, 65, 57, 50, 68, 67, 68, 68...
#> $ n_intervals <int> 176, 73, 73, 183, 46, 177, 111, 173, 149, 121, 182, 18...
#> $ n_matches  <int> 106, 36, 49, 157, 25, 149, 65, 130, 103, 73, 162, 138...
#> $ n_misses   <int> 70, 37, 24, 26, 21, 28, 46, 43, 46, 48, 20, 42, 25, 39...
#> $ prop_matches <dbl> 0.602, 0.493, 0.671, 0.858, 0.543, 0.842, 0.586, 0.751...
```

We have aggregated data with one row per child per aligner per sound class. age is provided in months or rescaled age_scale. Aggregation statistics include the number of files that were aligned by all 5 alignment algorithms (n_files), the number of intervals across the files (n_intervals), the number of matches and misses (n_matches, n_misses) and the proportion of matches (prop_matches).

Let's start with a logistic regression model with fixed effects for phoneme class, aligner, and class-by-aligner interactions and random by-child random intercepts and by-child-by-aligner random intercepts.

```
library(lme4)
#> Warning: package 'lme4' was built under R version 4.0.3
#> Loading required package: Matrix
#>
#> Attaching package: 'Matrix'
#> The following objects are masked from 'package:tidyr':
#>
#>     expand, pack, unpack

# Sound class x aligner interactions
m_class_int <- glmer(
  cbind(n_matches, n_misses) ~ class * aligner + (1 | child/aligner),
  family = binomial,
  model_data
)
```

By-child intercepts allow children to vary in alignment accuracy (some children will be easier or more difficult across all the aligners). By-child-by-aligner intercepts allows the alignment accuracy to vary by aligner and by child (some children will be easier or harder to align for particular aligners). Alternative random effects structures are plausible, but this structure is simple (matches are nested in children and aligner errors are nested in children) and yields model convergence.

```
summary(m_class_int)
#> Generalized linear mixed model fit by maximum likelihood (Laplace
#> Approximation) [glmerMod]
#> Family: binomial ( logit )
#> Formula: cbind(n_matches, n_misses) ~ class * aligner + (1 | child/aligner)
#> Data: model_data
#>
#>      AIC      BIC   logLik deviance df.resid
#>  6847.6   6951.8  -3401.8   6803.6     818
#>
#> Scaled residuals:
#>      Min       1Q   Median       3Q      Max
#> -4.9270 -0.9098  0.0501  0.9505  4.6142
#>
#> Random effects:
#> Groups          Name          Variance Std.Dev.
#> aligner:child (Intercept) 0.04887  0.2211
#> child          (Intercept) 0.05413  0.2326
#> Number of obs: 840, groups: aligner:child, 210; child, 42
#>
#> Fixed effects:
#>
#>              Estimate Std. Error z value Pr(>|z|)
#> (Intercept)      1.121585    0.057658  19.452 < 2e-16 ***
#> classothers     -0.198738    0.043431  -4.576 4.74e-06 ***
#> classplosives   -0.003951    0.038144  -0.104 0.917504
#> classvowels      0.364351    0.037020   9.842 < 2e-16 ***
#> alignerMFA (SAT)  0.739373    0.067438  10.964 < 2e-16 ***
#> alignerMFA (No SAT) -0.014103    0.063718  -0.221 0.824832
#> alignerP2FA     -0.643346    0.062301 -10.326 < 2e-16 ***
#> alignerProsodyLab -0.738173    0.062237 -11.861 < 2e-16 ***
#> classothers:alignerMFA (SAT) -0.241038    0.067174  -3.588 0.000333 ***
#> classplosives:alignerMFA (SAT) -0.097876    0.060341  -1.622 0.104796
#> classvowels:alignerMFA (SAT)  0.020632    0.059712   0.346 0.729695
#> classothers:alignerMFA (No SAT) -0.061060    0.061038  -1.000 0.317136
#> classplosives:alignerMFA (No SAT) -0.020203    0.053747  -0.376 0.707005
#> classvowels:alignerMFA (No SAT)  0.401899    0.053735   7.479 7.47e-14 ***
#> classothers:alignerP2FA      0.474701    0.059262   8.010 1.14e-15 ***
#> classplosives:alignerP2FA     0.018713    0.050961   0.367 0.713466
#> classvowels:alignerP2FA      0.322243    0.049697   6.484 8.92e-11 ***
#> classothers:alignerProsodyLab -0.002868    0.058233  -0.049 0.960725
#> classplosives:alignerProsodyLab -0.371961    0.050626  -7.347 2.02e-13 ***
#> classvowels:alignerProsodyLab  0.463956    0.049714   9.332 < 2e-16 ***
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#>
#> Correlation matrix not shown by default, as p = 20 > 12.
```

```
#> Use print(x, correlation=TRUE) or  
#>      vcov(x)      if you need it
```

We do not worry about interpreting these coefficients. They are on the log-odds scale and involve interactions of two dummy-coded categorical variables. We instead look at the estimated marginal (cell) means using the emmeans package.

First, we have to prepare a variance estimate for the bias-adjustment that's used by emmeans. A description of bias-adjustment procedure is available online <https://cran.r-project.org/web/packages/emmeans/vignettes/transformations.html#cbpp>.

```
variances <- broom.mixed::tidy(m_class_int, "ran_pars", scales = "vcov")  
#> Registered S3 method overwritten by 'broom.mixed':  
#>   method      from  
#> tidy.gamlss broom  
sigma <- sqrt(sum(variances$estimate))
```

Question 1 (Accuracy by aligner)

We get the marginal means on the probability scale (type = "response") for each aligner (~aligner).

```
library(emmeans)  
#> Warning: package 'emmeans' was built under R version 4.0.3  
  
means_by_aligner <- emmeans(  
  m_class_int,  
  spec = ~ aligner,  
  type = "response",  
  bias.adjust = TRUE,  
  sigma = sigma  
)  
#> NOTE: Results may be misleading due to involvement in interactions
```

Our model used included aligner-by-class interactions, but we are ignoring them for this estimate and averaging over the sound classes. The software prints **NOTE** to remind of this fact. Here are the means:

```
means_by_aligner  
#> aligner      prob      SE df asymp.LCL asymp.UCL  
#> Kaldi        0.757 0.00932 Inf      0.738      0.775  
#> MFA (SAT)    0.856 0.00638 Inf      0.843      0.868  
#> MFA (No SAT) 0.769 0.00902 Inf      0.750      0.786  
#> P2FA         0.669 0.01109 Inf      0.647      0.691  
#> ProsodyLab   0.607 0.01191 Inf      0.583      0.630  
#>  
#> Results are averaged over the levels of: class  
#> Confidence Level used: 0.95  
#> Intervals are back-transformed from the Logit scale  
#> Bias adjustment applied based on sigma = 0.32093
```

The **prob** is the estimated probability of a matching alignment, and **asymp.LCL** and **asymp.UCL** is a 95% interval on that estimated probability. These marginal means are pretty close to the means of the child-level proportions:

```
model_data %>%
  group_by(aligner) %>%
  summarise(
    empirical_proportion = mean(prop_matches),
    .groups = "drop"
  ) %>%
  left_join(
    broom::tidy(means_by_aligner, conf.int = TRUE),
    by = c("aligner")
  ) %>%
  select(
    aligner,
    empirical_proportion,
    marginal_probability = prob,
    lower_ci = asymp.LCL,
    upper_ci = asymp.UCL
  )
#> # A tibble: 5 x 5
#>   aligner      empirical_proportion marginal_probability lower_ci upper_ci
#>   <chr>          <dbl>          <dbl>          <dbl>    <dbl>
#> 1 Kaldi          0.752          0.757          0.738    0.775
#> 2 MFA (SAT)      0.854          0.856          0.843    0.868
#> 3 MFA (No SAT)  0.762          0.769          0.750    0.786
#> 4 P2FA          0.667          0.669          0.647    0.691
#> 5 ProsodyLab    0.599          0.607          0.583    0.630
```

Now, we can compare means to each other. We look at all pairwise differences between aligners (`pairwise ~ aligner`). `emmeans` handles the contrast-coding and the *p*-value adjustment automatically. Here we use the Bonferroni method.

```
by_aligner <- emmeans(
  m_class_int,
  spec = pairwise ~ aligner,
  type = "response",
  bias.adjust = TRUE,
  sigma = sigma,
  adjust = "bonferroni"
)
#> NOTE: Results may be misleading due to involvement in interactions
by_aligner$contrasts
#>   contrast      odds.ratio      SE df z.ratio p.value
#> Kaldi / MFA (SAT)      0.544 0.0290 Inf -12.386 <.0001
#> Kaldi / MFA (No SAT)  0.984 0.0516 Inf  -1.261 1.0000
#> Kaldi / P2FA          1.632 0.0849 Inf   8.447 <.0001
#> Kaldi / ProsodyLab    2.151 0.1118 Inf  13.779 <.0001
#> MFA (SAT) / MFA (No SAT) 1.904 0.1015 Inf  11.136 <.0001
#> MFA (SAT) / P2FA       3.156 0.1672 Inf  20.756 <.0001
#> MFA (SAT) / ProsodyLab  4.162 0.2201 Inf  26.012 <.0001
#> MFA (No SAT) / P2FA    1.743 0.0908 Inf   9.708 <.0001
#> MFA (No SAT) / ProsodyLab 2.298 0.1195 Inf  15.037 <.0001
#> P2FA / ProsodyLab     1.386 0.0716 Inf   5.354 <.0001
#>
#> Results are averaged over the levels of: class
#> P value adjustment: bonferroni method for 10 tests
```

```
#> Tests are performed on the log odds ratio scale  
#> Bias adjustment applied based on sigma = 0.32093
```

Differences here are in odds ratios like X / Y . Values greater than 1 indicate an advantage for X over Y, values less than 1 indicate an advantage for Y over X, and values of 1 indicates no difference in odds. Notably, there is not a significance difference between Kaldi and MFA (No SAT). These two are the most similar aligners in the comparison, a priori, because MFA is built on top of Kaldi and uses the same acoustic model behind the scenes.

Question 2 (Accuracy by sound class)

For question 2, we change from aggregating by aligner (averaging over class) to aggregating by class (averaging over aligner). We otherwise use the same procedure.

```
by_class <- emmeans(  
  m_class_int,  
  spec = pairwise ~ class,  
  type = "response",  
  bias.adjust = TRUE,  
  sigma = sigma,  
  adjust = "bonferroni"  
)  
#> NOTE: Results may be misleading due to involvement in interactions  
by_class$emmeans  
#>   class      prob      SE df asymp.LCL asymp.UCL  
#> fricatives 0.724 0.00806 Inf    0.708    0.740  
#> others     0.691 0.00869 Inf    0.674    0.708  
#> plosives   0.705 0.00826 Inf    0.688    0.721  
#> vowels     0.827 0.00573 Inf    0.815    0.838  
#>  
#> Results are averaged over the levels of: aligner  
#> Confidence level used: 0.95  
#> Intervals are back-transformed from the logit scale  
#> Bias adjustment applied based on sigma = 0.32093  
by_class$contrasts  
#>   contrast      odds.ratio      SE df z.ratio p.value  
#> fricatives / others      1.240 0.02415 Inf    8.461 <.0001  
#> fricatives / plosives     1.160 0.01986 Inf    5.738 <.0001  
#> fricatives / vowels       0.574 0.00979 Inf   -35.509 <.0001  
#> others / plosives         0.984 0.01774 Inf    -3.691 0.0013  
#> others / vowels          0.486 0.00875 Inf   -42.847 <.0001  
#> plosives / vowels         0.520 0.00801 Inf   -45.721 <.0001  
#>  
#> Results are averaged over the levels of: aligner  
#> P value adjustment: bonferroni method for 6 tests  
#> Tests are performed on the log odds ratio scale  
#> Bias adjustment applied based on sigma = 0.32093
```

Here the vowels have a consistent advantage over the other classes.

For question 2, we also looked at all of class-by-aligner effects. We use effect coding (`eff ~ ...`) to compare each phone-by-class mean to the average of the other phone-by-class

means. Again, emmeans handles the contrast and the *p*-value adjustment automatically. Here it uses "fdr" adjustment (false discovery rate, Benjamini & Hochberg, 1995).

```
by_aligner_class <- emmeans(
  m_class_int,
  spec = eff ~ class * aligner,
  type = "response",
  bias.adjust = TRUE,
  sigma = sigma
)
```

First, here are the means.

```
by_aligner_class$emmeans
#>   class aligner      prob      SE df asymp.LCL asymp.UCL
#>   fricatives Kaldi      0.749 0.01062 Inf      0.728      0.770
#>   others Kaldi      0.711 0.01191 Inf      0.687      0.734
#>   plosives Kaldi      0.749 0.01023 Inf      0.728      0.768
#>   vowels Kaldi      0.811 0.00827 Inf      0.794      0.826
#>   fricatives MFA (SAT)      0.861 0.00729 Inf      0.846      0.875
#>   others MFA (SAT)      0.801 0.00968 Inf      0.781      0.819
#>   plosives MFA (SAT)      0.849 0.00734 Inf      0.834      0.862
#>   vowels MFA (SAT)      0.901 0.00514 Inf      0.890      0.910
#>   fricatives MFA (No SAT)      0.747 0.01068 Inf      0.725      0.767
#>   others MFA (No SAT)      0.696 0.01220 Inf      0.671      0.719
#>   plosives MFA (No SAT)      0.742 0.01037 Inf      0.721      0.762
#>   vowels MFA (No SAT)      0.863 0.00657 Inf      0.849      0.875
#>   fricatives P2FA      0.614 0.01294 Inf      0.589      0.640
#>   others P2FA      0.676 0.01254 Inf      0.651      0.700
#>   plosives P2FA      0.618 0.01249 Inf      0.593      0.642
#>   vowels P2FA      0.757 0.00971 Inf      0.738      0.776
#>   fricatives Prosodylab      0.592 0.01317 Inf      0.566      0.618
#>   others Prosodylab      0.544 0.01392 Inf      0.517      0.571
#>   plosives Prosodylab      0.502 0.01315 Inf      0.476      0.528
#>   vowels Prosodylab      0.766 0.00951 Inf      0.747      0.784
#>
#> Confidence Level used: 0.95
#> Intervals are back-transformed from the logit scale
#> Bias adjustment applied based on sigma = 0.32093
```

More importantly, here are the contrasts.

```
by_aligner_class$contrasts
#>   contrast odds.ratio      SE df z.ratio p.value
#>   fricatives Kaldi effect      1.100 0.0461 Inf      1.087 0.3193
#>   others Kaldi effect      0.902 0.0395 Inf     -3.501 0.0006
#>   plosives Kaldi effect      1.096 0.0428 Inf      1.064 0.3193
#>   vowels Kaldi effect      1.584 0.0604 Inf     10.757 <.0001
#>   fricatives MFA (SAT) effect      2.305 0.1076 Inf     16.816 <.0001
#>   others MFA (SAT) effect      1.485 0.0692 Inf      7.405 <.0001
#>   plosives MFA (SAT) effect      2.082 0.0876 Inf     16.235 <.0001
#>   vowels MFA (SAT) effect      3.387 0.1431 Inf     27.699 <.0001
#>   fricatives MFA (No SAT) effect      1.085 0.0453 Inf      0.751 0.4762
#>   others MFA (No SAT) effect      0.837 0.0363 Inf     -5.265 <.0001
#>   plosives MFA (No SAT) effect      1.059 0.0412 Inf      0.186 0.8527
```

```
#> vowels MFA (No SAT) effect          2.335 0.0931 Inf  20.001 <.0001
#> fricatives P2FA effect              0.578 0.0230 Inf -15.025 <.0001
#> others P2FA effect                  0.762 0.0327 Inf  -7.491 <.0001
#> plosives P2FA effect                0.587 0.0220 Inf -15.546 <.0001
#> vowels P2FA effect                  1.149 0.0425 Inf   2.397 0.0207
#> fricatives Prosodylab effect         0.526 0.0209 Inf -17.450 <.0001
#> others Prosodylab effect             0.430 0.0180 Inf -21.395 <.0001
#> plosives Prosodylab effect           0.361 0.0134 Inf -28.709 <.0001
#> vowels Prosodylab effect             1.204 0.0448 Inf   3.648 0.0004
#>
#> P value adjustment: fdr method for 20 tests
#> Tests are performed on the log odds ratio scale
#> Bias adjustment applied based on sigma = 0.32093
```

In the manuscript, we report the classes for each aligner with the largest and smallest effects. Vowels have the largest effect for each aligner. The smallest effects are other sounds for Kaldi and the MFA aligners, plosives and fricatives for P2FA and plosives for Prosodylab.

Question 3 (Age effects)

We now augment the model to include age. We consider a model with just 2-way interactions and model with a full 3-way interactions.

```
# Add age, age x aligner, age x class effects
m_class_int_age_2way <- glmer(
  cbind(n_matches, n_misses) ~
    class * aligner + age_scale +
    age_scale:class + aligner:age_scale +
    (1 | child/aligner),
  family = binomial,
  model_data
)

# Add full 3-way interaction
m_class_int_age_3way <- glmer(
  cbind(n_matches, n_misses) ~
    class * aligner * age_scale +
    (1 | child/aligner),
  family = binomial,
  model_data
)

#> Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
#> Model failed to converge with max|grad| = 0.00225522 (tol = 0.002, component 1)
```

We disregard the convergence warning here because the diagnostic $\max|\text{grad}| = 0.00225522$ is barely above the threshold $\text{tol} = 0.002$.

Model comparison (lowest AIC, chi-square tests) supports the full, 3-way model.

```
format(
  anova(m_class_int, m_class_int_age_2way, m_class_int_age_3way),
  digits = 3
)
```

```
#>               npar  AIC  BIC LogLik deviance Chisq Df Pr(>Chisq)
#> m_class_int      22 6848 6952  -3402    6804    NA NA      NA
#> m_class_int_age_2way 30 6555 6697  -3248    6495 308.3  8  7.20e-62
#> m_class_int_age_3way 42 6520 6719  -3218    6436  59.2 12  3.11e-08
```

We want to compare the age-slopes between the aligners. `emtrends()` will compute the age trend (`var = "age_scale"`) for each aligner using effect coding (`eff ~ aligner`). We compare the slopes at 60 months (`at = list(age_scale = 0)`).

```
variances2 <- broom.mixed::tidy(
  m_class_int_age_3way,
  "ran_pars",
  scales = "vcov"
)
sigma2 <- sqrt(sum(variances2$estimate))

aligner_age_trends <- emtrends(
  m_class_int_age_3way,
  eff ~ aligner,
  var = "age_scale",
  bias.adjust = TRUE,
  sigma = sigma2,
  at = list(age_scale = 0)
)
#> NOTE: Results may be misleading due to involvement in interactions
aligner_age_trends
#> $emtrends
#> aligner      age_scale.trend      SE df asymp.LCL asymp.UCL
#> Kaldi          0.2326 0.0418 Inf    0.1506    0.3146
#> MFA (SAT)       0.0319 0.0430 Inf   -0.0523    0.1161
#> MFA (No SAT)    0.1731 0.0418 Inf    0.0910    0.2551
#> P2FA           0.1746 0.0414 Inf    0.0934    0.2557
#> Prosodylab     -0.0357 0.0413 Inf   -0.1168    0.0453
#>
#> Results are averaged over the levels of: class
#> Confidence level used: 0.95
#>
#> $contrasts
#> contrast      estimate      SE df z.ratio p.value
#> Kaldi effect    0.1173 0.0259 Inf   4.533 <.0001
#> MFA (SAT) effect -0.0834 0.0270 Inf  -3.094  0.0033
#> MFA (No SAT) effect 0.0578 0.0259 Inf   2.232  0.0256
#> P2FA effect     0.0593 0.0255 Inf   2.327  0.0249
#> Prosodylab effect -0.1510 0.0254 Inf  -5.945 <.0001
#>
#> Results are averaged over the levels of: class
#> P value adjustment: fdr method for 5 tests
```

The values here are on the logit (log-odds) scale. Exponentiating them produces odds ratios.

```
aligner_age_trends$emtrends %>%
  broom::tidy(conf.int = TRUE) %>%
  mutate(
    or_trend = exp(age_scale.trend),
```

```

    or_lcl = exp(asymp.LCL),
    or_ucl = exp(asymp.UCL),
    trend = age_scale.trend,
    se = std.error,
    z = z.ratio,
    p = p.value
  ) %>%
  select(aligner, starts_with("or_"), trend, se, z, p)
#> # A tibble: 5 x 8
#>   aligner      or_trend or_lcl or_ucl   trend     se      z      p
#>   <chr>      <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
#> 1 Kaldi      1.26  1.16  1.37  0.233  0.0418  5.56  0.000000269
#> 2 MFA (SAT)   1.03  0.949  1.12  0.0319  0.0430  0.743  0.458
#> 3 MFA (No SAT) 1.19  1.10  1.29  0.173  0.0418  4.14  0.0000354
#> 4 P2FA       1.19  1.10  1.29  0.175  0.0414  4.22  0.0000249
#> 5 ProsodyLab  0.965  0.890  1.05 -0.0357  0.0413 -0.865  0.387

```

Both the MFA-SAT and Prosodylab 95% intervals overlap with 1 (odds ratio) so we conclude that there is not a statistically clear age trend.

To describe the effect size on the response scale, we compare the match probabilities for the aligner with the steepest age effect (Kaldi). We calculate the match probability at age 0 and age 1 (`at = list(age_scale = c(0, 1))`) and then for each aligner compute pairwise differences between these probabilities to get the age 1 advantage over age 0 (`revpairwise ~ age_scale | aligner` computed reverse-pairwise differences—age 1 / age 0 on the odds ratio scale—but we do not use them here.)

```

aligner_age_trends_1_vs_0 <- emmeans(
  m_class_int_age_3way,
  revpairwise ~ age_scale | aligner ,
  bias.adjust = TRUE,
  sigma = sigma2,
  type = "response",
  at = list(age_scale = c(0, 1))
)
#> NOTE: Results may be misleading due to involvement in interactions

aligner_age_trends_1_vs_0$emmeans %>%
  broom::tidy(conf.int = TRUE) %>%
  filter(aligner == "Kaldi") %>%
  select(age_scale, aligner, prob, asymp.LCL, asymp.UCL)
#> # A tibble: 2 x 5
#>   age_scale aligner  prob asymp.LCL asymp.UCL
#>   <dbl> <chr>    <dbl>    <dbl>    <dbl>
#> 1      0 Kaldi   0.758    0.741    0.773
#> 2      1 Kaldi   0.797    0.777    0.816

```

Now, we can do the converse aggregation and averaging: Look at the age trend for each sound class averaging over aligners.

```

class_age_trends <- emtrends(
  m_class_int_age_3way,
  eff ~ class,
  var = "age_scale",

```

```

bias.adjust = TRUE,
sigma = sigma2,
at = list(age_scale = 0)
)
#> NOTE: Results may be misleading due to involvement in interactions
class_age_trends
#> $emtrends
#>   class      age_scale.trend      SE df asymp.LCL asymp.UCL
#> fricatives      0.25805 0.0349 Inf    0.18970    0.3264
#> others          0.13961 0.0353 Inf    0.07042    0.2088
#> plosives        0.05992 0.0340 Inf   -0.00671    0.1266
#> vowels          0.00356 0.0341 Inf   -0.06318    0.0703
#>
#> Results are averaged over the levels of: aligner
#> Confidence level used: 0.95
#>
#> $contrasts
#>   contrast      estimate      SE df z.ratio p.value
#> fricatives effect  0.1428 0.01112 Inf   12.844 <.0001
#> others effect      0.0243 0.01174 Inf    2.072 0.0383
#> plosives effect   -0.0554 0.00968 Inf   -5.721 <.0001
#> vowels effect     -0.1117 0.00973 Inf  -11.483 <.0001
#>
#> Results are averaged over the levels of: aligner
#> P value adjustment: fdr method for 4 tests

```

And likewise, look at odds ratios.

```

class_age_trends$emtrends %>%
  broom::tidy(conf.int = TRUE) %>%
  mutate(
    or_trend = exp(age_scale.trend),
    or_lcl = exp(asymp.LCL),
    or_ucl = exp(asymp.UCL),
    trend = age_scale.trend,
    se = std.error,
    z = z.ratio,
    p = p.value
  ) %>%
  select(class, starts_with("or_"), trend, se, z, p)
#> # A tibble: 4 x 8
#>   class      or_trend or_lcl or_ucl  trend      se      z      p
#>   <chr>      <dbl>  <dbl> <dbl>  <dbl>  <dbl> <dbl>  <dbl>
#> 1 fricatives    1.29  1.21   1.39 0.258  0.0349  7.40 1.36e-13
#> 2 others        1.15  1.07   1.23 0.140  0.0353  3.95 7.66e- 5
#> 3 plosives      1.06  0.993  1.13 0.0599 0.0340  1.76 7.80e- 2
#> 4 vowels        1.00  0.939  1.07 0.00356 0.0341  0.105 9.17e- 1

```

Finally, we can examine all of the Aligner x Class age trends.

```

class_aligner_trends <- emtrends(
  m_class_int_age_3way,
  ~ class | aligner,
  var = "age_scale",
  bias.adjust = TRUE,

```

```

sigma = sigma2,
type = "response",
at = list(age_scale = c(0, 1)),
)

class_aligner_trends %>%
  broom::tidy(conf.int = TRUE) %>%
  mutate(
    or_trend = exp(age_scale.trend),
    or_lcl = exp(asymp.LCL),
    or_ucl = exp(asymp.UCL),
    trend = age_scale.trend,
    se = std.error,
    z = z.ratio,
    p = p.value
  ) %>%
  select(aligner, class, starts_with("or_"), trend, se, z, p)
#> # A tibble: 20 x 9
#>   aligner      class or_trend or_lcl or_ucl trend      se      z      p
#>   <chr>      <chr>      <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
#> 1 Kaldi      fricativ~  1.46  1.32  1.61  0.376  0.0494  7.62  2.46e-14
#> 2 Kaldi      others    1.23  1.11  1.36  0.207  0.0509  4.07  4.64e- 5
#> 3 Kaldi      plosives  1.25  1.14  1.37  0.224  0.0462  4.84  1.32e- 6
#> 4 Kaldi      vowels    1.13  1.03  1.24  0.123  0.0455  2.71  6.76e- 3
#> 5 MFA (SAT)   fricativ~  1.30  1.17  1.45  0.264  0.0536  4.93  8.10e- 7
#> 6 MFA (SAT)   others    1.01  0.914  1.13  0.0146  0.0535  0.273  7.84e- 1
#> 7 MFA (SAT)   plosives  0.991  0.900  1.09  -0.00925  0.0489  -0.189  8.50e- 1
#> 8 MFA (SAT)   vowels    0.867  0.787  0.956  -0.142  0.0496  -2.87  4.17e- 3
#> 9 MFA (No SAT) fricativ~  1.37  1.25  1.51  0.317  0.0490  6.47  9.82e-11
#> 10 MFA (No SAT) others    1.19  1.08  1.31  0.174  0.0503  3.46  5.44e- 4
#> 11 MFA (No SAT) plosives  1.13  1.03  1.24  0.123  0.0459  2.67  7.49e- 3
#> 12 MFA (No SAT) vowels    1.08  0.986  1.19  0.0784  0.0470  1.67  9.54e- 2
#> 13 P2FA      fricativ~  1.29  1.18  1.42  0.258  0.0471  5.47  4.39e- 8
#> 14 P2FA      others    1.23  1.12  1.36  0.211  0.0500  4.21  2.50e- 5
#> 15 P2FA      plosives  1.09  1.00  1.19  0.0883  0.0447  1.98  4.82e- 2
#> 16 P2FA      vowels    1.15  1.06  1.26  0.141  0.0444  3.18  1.46e- 3
#> 17 Prosodylab fricativ~  1.08  0.983  1.18  0.0746  0.0468  1.59  1.11e- 1
#> 18 Prosodylab others    1.10  0.996  1.21  0.0913  0.0489  1.87  6.16e- 2
#> 19 Prosodylab plosives  0.882  0.808  0.962  -0.126  0.0445  -2.83  4.73e- 3
#> 20 Prosodylab vowels    0.833  0.763  0.909  -0.183  0.0446  -4.11  4.04e- 5

```

Computing environment

```

sessioninfo::session_info()
#> - Session info -----
#> setting value
#> version R version 4.0.2 (2020-06-22)
#> os Windows 10 x64
#> system x86_64, mingw32
#> ui RTerm
#> language (EN)
#> collate English_United States.1252
#> ctype English_United States.1252
#> tz America/Chicago
#> date 2020-10-30

```

```
#>
#> - Packages -----
#> ! package      * version date      lib source
#> assertthat    0.2.1  2019-03-21 [1] CRAN (R 4.0.2)
#> backports     1.1.10 2020-09-15 [1] CRAN (R 4.0.2)
#> base64url     1.4    2018-05-14 [1] CRAN (R 4.0.2)
#> blob          1.2.1  2020-01-20 [1] CRAN (R 4.0.2)
#> boot          1.3-25 2020-04-26 [1] CRAN (R 4.0.2)
#> broom         0.7.2  2020-10-20 [1] CRAN (R 4.0.2)
#> broom.mixed   0.2.6  2020-05-17 [1] CRAN (R 4.0.2)
#> cellranger    1.1.0  2016-07-27 [1] CRAN (R 4.0.2)
#> cli           2.1.0  2020-10-12 [1] CRAN (R 4.0.2)
#> coda          0.19-4 2020-09-30 [1] CRAN (R 4.0.2)
#> codetools     0.2-16 2018-12-24 [1] CRAN (R 4.0.0)
#> colorspace    1.4-1  2019-03-18 [1] CRAN (R 4.0.2)
#> crayon        1.3.4  2017-09-16 [1] CRAN (R 4.0.2)
#> DBI           1.1.0  2019-12-15 [1] CRAN (R 4.0.2)
#> dbplyr        1.4.4  2020-05-27 [1] CRAN (R 4.0.2)
#> digest        0.6.27 2020-10-24 [1] CRAN (R 4.0.3)
#> dplyr         * 1.0.2  2020-08-18 [1] CRAN (R 4.0.2)
#> drake         7.12.6 2020-10-10 [1] CRAN (R 4.0.3)
#> ellipsis      0.3.1  2020-05-15 [1] CRAN (R 4.0.2)
#> emmeans       * 1.5.2-1 2020-10-25 [1] CRAN (R 4.0.3)
#> estimability  1.3    2018-02-11 [1] CRAN (R 4.0.0)
#> evaluate      0.14   2019-05-28 [1] CRAN (R 4.0.2)
#> fansi         0.4.1  2020-01-08 [1] CRAN (R 4.0.2)
#> filelock      1.0.2  2018-10-05 [1] CRAN (R 4.0.2)
#> forcats       * 0.5.0  2020-03-01 [1] CRAN (R 4.0.2)
#> fs            1.5.0  2020-07-31 [1] CRAN (R 4.0.2)
#> generics      0.0.2  2018-11-29 [1] CRAN (R 4.0.2)
#> ggplot2       * 3.3.2  2020-06-19 [1] CRAN (R 4.0.2)
#> glue          1.4.2  2020-08-27 [1] CRAN (R 4.0.2)
#> gtable        0.3.0  2019-03-25 [1] CRAN (R 4.0.2)
#> haven         2.3.1  2020-06-01 [1] CRAN (R 4.0.2)
#> hms           0.5.3  2020-01-08 [1] CRAN (R 4.0.2)
#> htmltools     0.5.0  2020-06-16 [1] CRAN (R 4.0.2)
#> http         1.4.2  2020-07-20 [1] CRAN (R 4.0.2)
#> igraph        1.2.6  2020-10-06 [1] CRAN (R 4.0.2)
#> jsonlite      1.7.1  2020-09-07 [1] CRAN (R 4.0.2)
#> knitr         1.30   2020-09-22 [1] CRAN (R 4.0.2)
#> Lattice       0.20-41 2020-04-02 [1] CRAN (R 4.0.2)
#> lifecycle     0.2.0  2020-03-06 [1] CRAN (R 4.0.2)
#> lme4          * 1.1-25 2020-10-23 [1] CRAN (R 4.0.3)
#> lubridate     1.7.9  2020-06-08 [1] CRAN (R 4.0.2)
#> magrittr      1.5    2014-11-22 [1] CRAN (R 4.0.2)
#> MASS          7.3-53 2020-09-09 [1] CRAN (R 4.0.2)
#> Matrix        * 1.2-18 2019-11-27 [1] CRAN (R 4.0.2)
#> minqa         1.2.4  2014-10-09 [1] CRAN (R 4.0.2)
#> modelr        0.1.8  2020-05-19 [1] CRAN (R 4.0.2)
#> multcomp      1.4-14 2020-09-23 [1] CRAN (R 4.0.2)
#> munsell       0.5.0  2018-06-12 [1] CRAN (R 4.0.2)
#> mvtnorm       1.1-1  2020-06-09 [1] CRAN (R 4.0.0)
#> nlme          3.1-150 2020-10-24 [1] CRAN (R 4.0.3)
#> nloptr        1.2.2.2 2020-07-02 [1] CRAN (R 4.0.2)
#> pillar        1.4.6  2020-07-10 [1] CRAN (R 4.0.2)
```

```
#> pkgconfig      2.0.3    2019-09-22 [1] CRAN (R 4.0.2)
#> plyr            1.8.6    2020-03-03 [1] CRAN (R 4.0.2)
#> prettyunits     1.1.1    2020-01-24 [1] CRAN (R 4.0.2)
#> progress        1.2.2    2019-05-16 [1] CRAN (R 4.0.2)
#> purrr           * 0.3.4    2020-04-17 [1] CRAN (R 4.0.2)
#> R6              2.5.0    2020-10-28 [1] CRAN (R 4.0.2)
#> Rcpp            1.0.5    2020-07-06 [1] CRAN (R 4.0.2)
#> readr           * 1.4.0    2020-10-05 [1] CRAN (R 4.0.2)
#> readxl          1.3.1    2019-03-13 [1] CRAN (R 4.0.2)
#> reprex          0.3.0    2019-05-16 [1] CRAN (R 4.0.2)
#> reshape2       1.4.4    2020-04-09 [1] CRAN (R 4.0.2)
#> rlang           0.4.8    2020-10-08 [1] CRAN (R 4.0.3)
#> rmarkdown       2.5      2020-10-21 [1] CRAN (R 4.0.2)
#> rstudioapi      0.11     2020-02-07 [1] CRAN (R 4.0.2)
#> rvest           0.3.6    2020-07-25 [1] CRAN (R 4.0.2)
#> sandwich        3.0-0    2020-10-02 [1] CRAN (R 4.0.2)
#> scales          1.1.1    2020-05-11 [1] CRAN (R 4.0.2)
#> sessioninfo     1.1.1    2018-11-05 [1] CRAN (R 4.0.2)
#> statmod         1.4.35   2020-10-19 [1] CRAN (R 4.0.3)
#> storr           1.2.4    2020-10-12 [1] CRAN (R 4.0.3)
#> stringi         1.5.3    2020-09-09 [1] CRAN (R 4.0.2)
#> stringr         * 1.4.0    2019-02-10 [1] CRAN (R 4.0.2)
#> survival        3.2-7    2020-09-28 [1] CRAN (R 4.0.2)
#> TH.data         1.0-10   2019-01-21 [1] CRAN (R 4.0.2)
#> tibble          * 3.0.4    2020-10-12 [1] CRAN (R 4.0.3)
#> tidyr           * 1.1.2    2020-08-27 [1] CRAN (R 4.0.2)
#> tidyselect      1.1.0    2020-05-11 [1] CRAN (R 4.0.2)
#> tidyverse       * 1.3.0    2019-11-21 [1] CRAN (R 4.0.2)
#> D TMB           1.7.18   2020-07-27 [1] CRAN (R 4.0.2)
#> txtq            0.2.3    2020-06-23 [1] CRAN (R 4.0.2)
#> utf8            1.1.4    2018-05-24 [1] CRAN (R 4.0.2)
#> vctrs           0.3.4    2020-08-29 [1] CRAN (R 4.0.2)
#> withr           2.3.0    2020-09-22 [1] CRAN (R 4.0.2)
#> xfun            0.18     2020-09-29 [1] CRAN (R 4.0.2)
#> xml2            1.3.2    2020-04-23 [1] CRAN (R 4.0.2)
#> xtable          1.8-4    2019-04-21 [1] CRAN (R 4.0.2)
#> yaml            2.2.1    2020-02-01 [1] CRAN (R 4.0.0)
#> zoo             1.8-8    2020-05-02 [1] CRAN (R 4.0.2)
#>
#> [1] C:/Users/Tristan/Documents/R/win-library/4.0
#> [2] C:/Program Files/R/R-4.0.2/Library
#>
#> D -- DLL MD5 mismatch, broken installation.
```

Benjamini, Y., and Hochberg, Y. (1995). Controlling the false discovery rate: a practical and powerful approach to multiple testing. *Journal of the Royal Statistical Society Series B*, 57, 289–300. <http://www.jstor.org/stable/2346101>.